

# **Estimating abundance from occupancy data: the Royle-Nichols model**

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Since (repeated) occupancy surveys allow us to estimate detection probability, this should also provide information about abundance.

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### WARNING

Results may be sensitive to assumptions that are difficult or impossible to verify.

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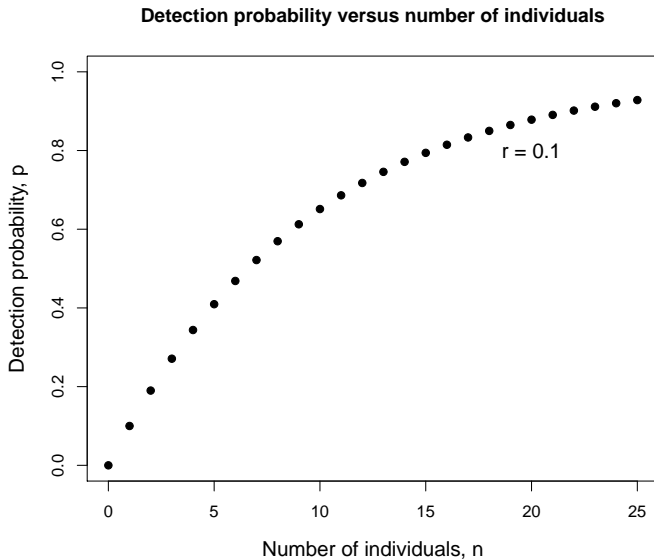
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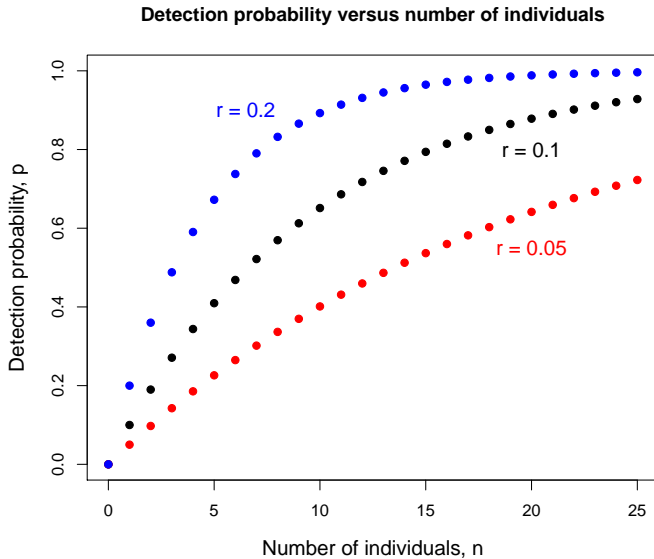
Then it follows that

$$p = 1 - (1 - r)^N.$$

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Given the value of  $N = n$  we can substitute

$$p = 1 - (1 - r)^n$$

to get a **conditional** probability

$$\Pr(W = w | N = n) = \text{function of } K, w, r \text{ and } n.$$

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Then we can work out the **unconditional** probability  $\Pr(W = w)$  using the law of total probability

$$\Pr(W = w) = \sum_{n=0}^{\infty} \Pr(W = w|N = n) \times \Pr(N = n).$$

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and the parameters to estimate are  $r$  and  $\mu$ .

## What happened to occupancy parameter?

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Answer: Still there, given that we have a probability distribution for  $N$  ...

$$\psi = 1 - \Pr(N = 0)$$

E.g. If  $N$  has Poisson distribution with mean  $\mu$

$$\psi = 1 - \exp(-\mu)$$

## Example: Wood thrush (from R package unmarked)

Data for wood thrush (*Hylocichla mustelina*) from North American Breeding Bird survey (1991)

Point counts at 50 locations along a survey route in New Hampshire collected over a 30 day period

11 sampling occasions (same observer)

Count	0	1	2	3	4
Frequency	344	162	40	3	1

Counts not always reliable, so reduce data to detection/no detection

## Fitting the basic Royle-Nichols model

Model with constant abundance ( $\mu$ ) and  
constant detection probability ( $r$ )

AIC: 633.9534

Estimate of  $\log(\mu)$

Abundance:

Estimate	SE	z	P(> z )
0.792	0.158	5.03	5e-07

Estimate of  $\text{logit}(r)$

Detection:

Estimate	SE	z	P(> z )
-1.21	0.17	-7.14	9.41e-13



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Estimate of  $\text{logit}(r)$              $[\hat{r} = 1/(1 + \exp(1.21)) = 0.23]$

Detection:

Estimate	SE	z	P(> z )
-1.21	0.17	-7.14	9.41e-13

## Fitting the basic Royle-Nichols model - implications

Number of individuals at a site assumed to follow a Poisson distribution with mean estimated as

$$\hat{\mu} = 2.21$$

In particular, proportion of occupied sites ( $\psi$ ) estimated as

$$\Pr(N > 0) = 1 - \exp(-2.21) = 0.89$$

(Observed proportion of sites where species detected =  $45/50 = 0.9$ )

## Fitting the basic Royle-Nichols model - implications

Probability of detecting **an individual** estimated as

$$\hat{r} = 0.23$$

E.g. With 2 individuals per site, detection probability  
 $= 1 - (1 - r)^2 = 0.41$ .

Probability of failing to detect on any of 11 visits to a site with 2  
individuals  $= 1 - (1 - 0.41)^{11} = 0.003$ .

Almost certain to detect species on at least one occasion if present

## Empirical Bayes estimates of numbers at each site

Previously we had the conditional probability

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$$\Pr(N = n|W = w) = \frac{\Pr(W = w|N = n) \times \Pr(N = n)}{\sum_{k=0}^{\infty} \Pr(W = w|N = k) \times \Pr(N = k)}$$

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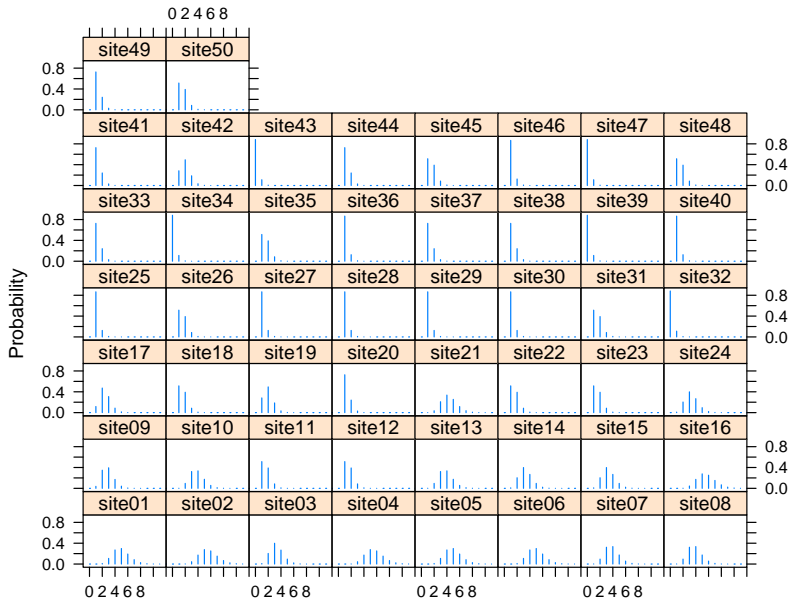
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Once we have fitted the model, we can estimate these probabilities for each site.

The resulting estimators of  $\Pr(N = n|W = w)$  are called [empirical Bayes](#) estimators.

# Empirical Bayes estimates at each site for Woodthrush data



## Aside: Back to the point counts

If  $N$  varies across sites as Poisson distribution with mean 2.21 and individuals are detected with probability 0.23, then expected distribution of point counts is Poisson with mean  $2.21 \times 0.23 = 0.51$ .



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Observed frequency	344	162	40	3	1
Expected frequency	331.2	168.0	42.6	7.2	1.0

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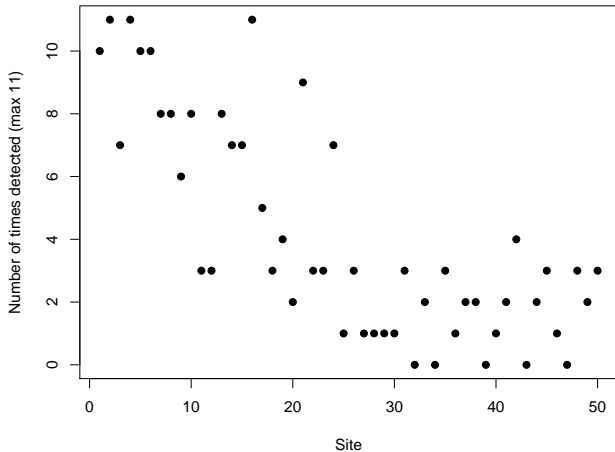
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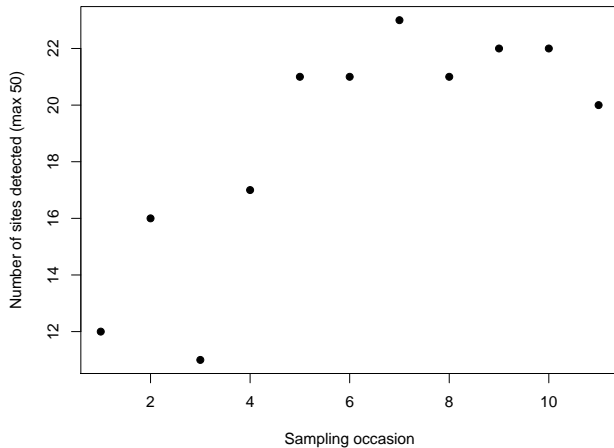
Quite a good fit (chi-squared = 3.32, 3 d.f.)

But ...

## Abundance related to site?



## Detection related to sampling visit?



## Fitting Royle-Nichols model with covariates

Model with abundance ( $\mu$ ) as a function of site number  
detection probability ( $r$ ) as a function of  
observation number

AIC: 562.0886 (constant model AIC: 633.9534)

Abundance:

	Estimate	SE	z	P(> z )
(Intercept)	2.9672	0.21465	13.82	1.85e-43
Site number	-0.0676	0.00778	-8.69	3.76e-18

Detection:

	Estimate	SE	z	P(> z )
(Intercept)	-2.5607	0.2641	-9.70	3.15e-22
Obs number	0.0791	0.0254	3.11	1.84e-03

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Other distributions possible in principle (e.g. negative binomial) but may be hard to fit in practice.

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Goodness of fit?